

# Admixture affects the rate and repeatability of experimental adaptation to a stressful environment in *Callosobruchus maculatus*

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## Abstract

Admixture is common in nature, and can serve as a crucial source of adaptive potential through the generation of novel genotype combinations and phenotypes. Conversely, the presence of hybrid incompatibilities can decrease the fitness of hybrids. Due to the pervasiveness of admixture in nature and its potential role in facilitating adaptation, understanding how admixture affects the rate and repeatability of evolution is important for furthering our understanding of evolutionary dynamics. However, few studies have assessed how patterns of evolutionary parallelism in admixed lineages are affected by the presence of strong ecological pressure. In this experiment, we assessed patterns of evolution and parallelism across admixed and non-admixed cowpea seed beetles (*Callosobruchus maculatus*) during adaptation to a novel, stressful host: lentil. Specifically, we asked (1) whether admixture facilitates adaptation to lentil, (2) whether parallelism was higher in admixed or non-admixed lineages, and (3) to what degree parallelism in admixed lineages was associated with selection on globally adaptive alleles versus epistatic effects and hybrid incompatibilities. We found that admixture facilitated adaptation to lentil, and evolutionary rescue—defined as adaptation that prevents population extinction—occurred in all lineages. The degree of evolutionary parallelism was highest in two admixed lineages, but notable in all lineages. Adaptation to lentil appeared to be driven by selection on alleles that were globally adaptive. However, even during evolutionary rescue in a marginal environment, the purging of hybrid incompatibilities appeared to contribute substantially to evolutionary parallelism in admixed lineages.

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# Abstract

Admixture is common in nature, and can serve as a crucial source of adaptive potential through the generation of novel genotype combinations and phenotypes. Conversely, the presence of hybrid incompatibilities can decrease the fitness of hybrids. Due to the pervasiveness of admixture in nature and its potential role in facilitating adaptation, understanding how admixture affects the rate and repeatability of evolution is important for furthering our understanding of evolutionary dynamics. However, few studies have assessed how patterns of evolutionary parallelism in admixed lineages are affected by the presence of strong ecological pressure. In this experiment, we assessed patterns of evolution and parallelism across admixed and non-admixed cowpea seed beetles (*Callosobruchus maculatus*) during adaptation to a novel, stressful host: lentil. Specifically, we asked (1) whether admixture facilitates adaptation to lentil, (2) whether parallelism was higher in admixed or non-admixed lineages, and (3) to what degree parallelism in admixed lineages was associated with selection on globally adaptive alleles versus epistatic effects and hybrid incompatibilities. We found that admixture facilitated adaptation to lentil, and evolutionary rescue—defined as adaptation that prevents population extinction—occurred in all lineages. The degree of evolutionary parallelism was highest in two admixed lineages, but notable in all lineages. Adaptation to lentil appeared to be driven by selection on alleles that were globally adaptive. However, even during evolutionary rescue in a marginal environment, the purging of hybrid incompatibilities appeared to contribute substantially to evolutionary parallelism in admixed lineages.

**Keywords:** *Callosobruchus maculatus*, adaptation, parallel evolution, Bayesian linear models, admixture, evolutionary rescue

## 23 Introduction

24 Admixture is increasingly being recognized as a major driver of evolutionary dynamics, as  
25 well as a potentially critical source of adaptive potential. Admixture is a widespread phe-  
26 nomenon, occurring in at least 10% of animal and 25% of plant species (Mallet, 2005), and a  
27 substantial portion of many species' genomes—including our own—are derived from hybrid  
28 origins (Gompert *et al.*, 2006; Hermansen *et al.*, 2011; Edwards *et al.*, 2011; Sankararaman  
29 *et al.*, 2016; Schumer *et al.*, 2016; Meier *et al.*, 2017; Short & Streisfeld, 2023; Rosser *et al.*,  
30 2024). Admixture events can result in the transfer of just a few alleles from one population  
31 to another (i.e. adaptive introgression) (Enard & Petrov, 2018; Oziolor *et al.*, 2019; Nanaei  
32 *et al.*, 2023; Rossi *et al.*, 2024), the reinforcement of species boundaries (Bewick & Dyer, 2014;  
33 Turissini & Matute, 2017; Bhargav *et al.*, 2022), or in some cases, genome stabilization and  
34 the formation of stable mosaic hybrid species (Gompert *et al.*, 2006; Mallet, 2007; Schumer  
35 *et al.*, 2018; Sun *et al.*, 2020; Rosser *et al.*, 2024). By bringing together new combinations  
36 of alleles from previously isolated parental populations, admixture can create novel pheno-  
37 typic variation (i.e. transgressive segregation) and serve as a source of evolutionary novelty  
38 (Lewontin & Birch, 1966; Rieseberg *et al.*, 1999; Pereira *et al.*, 2014; Chhina *et al.*, 2022).  
39 The extreme phenotypes generated by admixture combined with the transfer of globally  
40 beneficial alleles (i.e. adaptive introgression) and the genetic benefits of outbreeding (e.g.,  
41 heterosis and the masking of deleterious recessive alleles) can increase the adaptive potential  
42 of admixed populations, particularly in novel or marginal environments (Crow, 1948; Buerkle  
43 *et al.*, 2000; Gompert *et al.*, 2006; De Carvalho *et al.*, 2010; Oziolor *et al.*, 2019; Durkee *et al.*,  
44 2023). Conversely, the presence of Dobzhansky-Muller incompatibilities (Dobzhansky, 1982)  
45 and the breakdown of adaptive gene complexes can reduce fitness in admixed individuals  
46 (i.e. outbreeding depression), leading to selective pressure against hybridization (Verhoeven  
47 *et al.*, 2011; Turissini & Matute, 2017; Kim *et al.*, 2018; Calvo-Baltanás *et al.*, 2021; Bhargav  
48 *et al.*, 2022; Mantel & Sweigart, 2024). Because admixed populations are subject to multiple  
49 conflicting evolutionary pressures, the evolutionary outcomes of admixture vary widely. As

50 such, determining the degree to which evolution in admixed populations is repeatable—and  
51 therefore predictable—is of particular interest for understanding how deterministic processes  
52 (e.g., natural selection imposed by the environment) and constraints imposed by admixture  
53 interact to shape patterns of genomic change.

54         The degree of repeatability in genome evolution post-admixture depends on many  
55 factors, including demographic history, the degree of genetic divergence between parental  
56 populations, recombination landscapes across the genome, and how far from the phenotypic  
57 optimum each parental population is in the environment where admixture occurs (Schumer  
58 *et al.*, 2018; Moran *et al.*, 2021; McFarlane *et al.*, 2022; Langdon *et al.*, 2024; Owens *et al.*,  
59 2025). A few general principles have already emerged regarding the repeatability of evolution  
60 at a genomic level post-admixture, including the purging of ancestry derived from the minor  
61 parental population—the parental population that contributed the least amount of ancestry  
62 to the hybrid genome (Schumer *et al.*, 2018; Chaturvedi *et al.*, 2020; Moran *et al.*, 2021;  
63 Langdon *et al.*, 2022, 2024). When Dobzhansky-Muller incompatibilities are present or  
64 intermediate hybrid phenotypes are ecologically unsuitable for the environment, purging  
65 ancestry from the minor parent can be the most direct evolutionary route for adaptation in  
66 admixed populations (Langdon *et al.*, 2022). Purging of minor parent ancestry may even be  
67 repeatable across hybrids formed from different species pairs (Langdon *et al.*, 2022, 2024).  
68 Similarly, when one parental population has a lower effective population size than the other  
69 (i.e. island versus mainland populations, see Matute *et al.*, 2020), mildly deleterious alleles  
70 that accumulated and fixed in the smaller population via genetic drift can result in strong  
71 selection against ancestry from that population (Harris & Nielsen, 2016; Juric *et al.*, 2016).  
72 Selective pressure against this hybridization load can lead to purging of entire blocks of local  
73 ancestry inherited from the smaller, more inbred population, especially at sites with low  
74 recombination rates (Matute *et al.*, 2020; Nouhaud *et al.*, 2022).

75         However, while a considerable amount of work has been done to determine factors  
76 shaping the repeatability of evolution in admixed populations in an organism’s native habi-

77 tat (Rieseberg *et al.*, 2003; Schumer *et al.*, 2018; Chaturvedi *et al.*, 2020; Langdon *et al.*,  
78 2022; Nouhaud *et al.*, 2022; Langdon *et al.*, 2024; Owens *et al.*, 2025) or under benign lab-  
79 oratory conditions (Matute *et al.*, 2020), few studies explicitly address the impact of strong  
80 directional selection imposed by stressful ecological conditions on patterns of evolutionary  
81 repeatability in admixed populations. Given the strong potential for admixture to facil-  
82 itate adaptation and evolutionary rescue—defined as adaptation that prevents population  
83 extinction—under stressful environmental conditions via the expression of transgressive phe-  
84 notypes and transfer of globally adaptive alleles (Lewontin & Birch, 1966; Gompert *et al.*,  
85 2006; De Carvalho *et al.*, 2010; Pereira *et al.*, 2014; Stelkens *et al.*, 2014; Oziolor *et al.*, 2019;  
86 Vedder *et al.*, 2022; Durkee *et al.*, 2023), this remains a critical gap in our understanding  
87 of the predictability and repeatability of evolution in admixed populations. In the face of  
88 unprecedented anthropogenic change, determining how strong ecological selection alters the  
89 genomic consequences of admixture is also of critical relevance for determining the effect of  
90 admixture on adaptive potential. While intrinsic hybrid incompatibilities commonly drive  
91 patterns of repeatability during the evolution of admixed populations (Chaturvedi *et al.*,  
92 2020; Matute *et al.*, 2020; Langdon *et al.*, 2022; Nouhaud *et al.*, 2022; Owens *et al.*, 2025),  
93 the severe population bottlenecks that occur during evolutionary rescue could drastically  
94 increase the degree of stochasticity experienced during adaptation, potentially reducing re-  
95 peatability (see McFarlane *et al.*, 2022). Conversely, when populations begin far from the  
96 phenotypic optimum, rapid adaptation during evolutionary rescue may be initially driven  
97 by selection of just a few major-effect loci (rather than many small-effect loci) (Orr, 2005;  
98 Alexander *et al.*, 2014). Selection concentrated on a few loci during bouts of rapid adapta-  
99 tion could potentially increase the repeatability of evolution during evolutionary rescue, but  
100 how the influx of novel standing genetic variation plus intrinsic incompatibilities introduced  
101 via admixture might alter patterns of selection and change during evolutionary rescue is  
102 unclear.

103 In this study, we used experimental evolution to assess how admixture affects pat-

104 terns of evolutionary rescue and repeatability in cowpea seed beetles, *Callosobruchus mac-*  
105 *ulatus*, during adaptation to a novel, stressful host. *Callosobruchus maculatus* is a globally-  
106 distributed pest of stored legumes from the tribe Phaseoleae (e.g., mung bean, adzuki bean,  
107 and cowpea; Tuda *et al.*, 2006; Kéb  *et al.*, 2017). Because cowpea seed beetles have been  
108 associated with human crop stores for thousands of years and their larvae spend the en-  
109 tirety of their development within a single seed, laboratory conditions closely approximate  
110 the “natural” habitat of *C. maculatus* (Messina, 1991; Tuda *et al.*, 2014; Kéb  *et al.*, 2017).  
111 Populations from different geographic locations vary substantially in fitness traits, including  
112 larval competitiveness, body size, oviposition preference, and fecundity (Credland & Dick,  
113 1987; Messina, 1991, 1993; Messina *et al.*, 2018; Burc *et al.*, 2025). Lentil (*Lens culinaris*,  
114 tribe Fabeae) is a particularly poor host for *C. maculatus* (Messina *et al.*, 2009). Initial  
115 survival on lentil is often less than 3%, and experimental attempts to establish *C. maculatus*  
116 populations on lentil sometimes result in extinction (Messina *et al.*, 2009, 2020). Despite  
117 this, *C. maculatus* lineages on lentil that do not go extinct have been found to rapidly re-  
118 bound, with percent survival rising to over 80% within 20 generations (Messina *et al.*, 2009;  
119 R go *et al.*, 2019). Previous ecological studies have shown that admixture likely facilitates  
120 adaptation to lentil in the cowpea seed beetle (Messina *et al.*, 2020), and previous genomic  
121 studies have found a modest degree of parallelism at a genomic level across non-admixed  
122 lineages during adaptation to lentil (Gompert & Messina, 2016; R go *et al.*, 2019). However,  
123 to date no studies have assessed how both admixture and environmental stress combined  
124 affect the repeatability of genomic change during adaptation a novel, stressful host.

125 Here, we assessed how admixture affects ecological (demographic) and evolutionary  
126 dynamics and the degree of evolutionary parallelism (repeatability of genomic change) during  
127 adaptation to lentil in *C. maculatus*. Specifically we asked the following questions: (1) to  
128 what degree does admixture facilitate adaptation to lentil, (2) is evolution more repeatable  
129 in admixed or non-admixed lineages during evolutionary rescue, and (3) to what degree is  
130 repeatability during evolutionary rescue in admixed lineages driven by (a) a shared genetic

131 basis for adaptation to lentil across admixed and non-admixed lineages (i.e. selection on  
132 globally-adaptive alleles where the beneficial effects do not depend on genetic background)  
133 versus (b) a shared genetic basis for adaptation in admixed (but not non-admixed) lineages  
134 independent of host plant, which would suggest epistatic effects in hybrid lineages and the  
135 purging of hybrid incompatibilities?

## 136 **Materials and Methods**

### 137 **Experimental Design**

138 We used cowpea-adapted lineages of *Callosobruchus maculatus* from three different conti-  
139 nents for this experiment: Burkina Faso (Africa), Brazil (South America), and California  
140 (North America) (Fig. 1). These lineages all originally utilized cowpea, *Vigna unguiculata*,  
141 as their native host, have non-competitive larvae, and show low initial survival rates on lentil  
142 (Messina *et al.*, 2020). All lineages were obtained from Dr. Charles Fox at the University of  
143 Kentucky (Messina *et al.*, 2018), but were originally collected from infested cowpeas in the  
144 field or in markets across the world. Cultures from all three lineages were maintained con-  
145 tinuously in the laboratory on cowpea after their initial collection. The Burkina Faso (BF)  
146 lineage was collected from a field of cowpeas (*V. unguiculata*) in Ouagadougou, Burkina  
147 Faso by Dr. J. Huignard at the University of Tours in 1989 (Messina, 1993; Messina *et al.*,  
148 2018). The Brazil population (BZ) was collected from Campinas, Brazil in 1975 (Tran &  
149 Credland, 1995) and later maintained by Dr. Robert Smith at the University of Leicester  
150 (Guedes *et al.*, 2003; Dowling *et al.*, 2007a). The North American lineage was collected from  
151 California (CA) and later maintained by Dr. Peter Credland at the University of London  
152 (Tuda *et al.*, 2014; Dowling *et al.*, 2007b). All lineages are estimated to have been main-  
153 tained under standard laboratory conditions in excess of 300 generations at the time of our  
154 experiment.



155 We began our experiment with a single stock colony each from the BF, BZ, and  
156 CA lineages (3 jars total). These colonies had been maintained in the laboratory at Utah  
157 State University in excess of 100 generations at the time of this experiment. Colonies were  
158 kept in 2 L glass jars containing approximately 750 g of cowpeas. New generations were  
159 founded by transferring  $\sim$ 2000 newly-emerged adult beetles (estimated by volume using  
160 an insect aspirator) to fresh culture jars once every 25-30 days (hereafter referred to as  
161 “standard culture”). During this experiment, all colonies were housed at 27°C with a 14/10  
162 day cycle in one of two Percival incubators (both model No. I-36VL). Due to the large  
163 amount of metabolic water produced by growing beetle larvae, we installed a dehumidifier in  
164 each incubator to reduce humidity levels to between 15-50%. While under standard culture  
165 conditions only 2000 adult beetles per generation are transferred, each jar will produce far  
166 more than 2000 adult beetles each generation. This allows us to split any given jar of beetles  
167 into multiple daughter colonies each month without imposing a population bottleneck on  
168 the original colony. To found the colonies for our experiment, each month from our stock  
169 colonies we removed (1) 2000 adult beetles to found the next generation of the stock colony,  
170 (2) 2000 adult beetles to found the purebred control colonies for that month’s replicates, and  
171 (3) approximately 1000 pupae-containing beans to use for that month’s admixed replicates.

172 To form our admixed lineages, we produced true F1 hybrids with a founding popu-  
173 lation size of 1000 parental beetles each. To accomplish this, one to three days before the  
174 beetles’ expected emergence date, we randomly sampled 1000 pupae-containing beans from  
175 each of our parental stock colonies. These beans were isolated in individual cellulose pill  
176 capsules and maintained in the incubator under standard conditions. Twice daily, emerging  
177 virgin beetles were tallied and sorted into petri dishes by population and sex. This process  
178 was continued until we had collected four dishes of 250 virgin beetles each from each stock  
179 colony: two all-male and two all-female dishes from each parental population (BF, BZ, and  
180 CA). Reciprocal crosses were then performed for each combination of parental populations.  
181 In other words, we placed 250 virgin males from the first parental population in a jar with

182 250 virgin females from the second, and vice versa in a second jar. After 10 days (at which  
183 point most or all of the purebred adults had died), we combined each male×female jar with  
184 its reciprocal female×male pair to found a single admixed colony comprised of the true F1  
185 offspring of the 1000 purebred founding beetles. This method ensured we were producing  
186 admixed lineages with equal genetic contribution from both sexes from each parental pop-  
187 ulations. To found our purebred control colonies, we simply transferred 1000 beetles (as  
188 measured by volume) from each purebred stock colony to fresh culture jars. We produced  
189 11 full replicates, where each replicate consisted of three purebred (BF, BZ, CA) and three  
190 admixed (BF×BZ, BF×CA, BZ×CA) cowpea colonies each, for a total of 66 cowpea colonies  
191 with a founding population size of 1000 beetles per jar (Fig. 1).

192         After maintaining both our purebred and admixed colonies on their native host (cow-  
193 pea) for two successive generations post-admixture, we split each of our 66 cowpea colonies to  
194 form 66 additional colonies on our novel, stressful host: lentil. To do so, we removed a total  
195 of 4000 adult beetles (as measured by volume) from each our our 66 F2 cowpea colonies and  
196 transferred 2000 to fresh lentil culture jars and 2000 to fresh cowpea culture jars to lay eggs.  
197 Thus, the first generation of beetle larvae to feed on the novel food source in our experiment  
198 was the F3 generation. This left us with a total of 132 beetle colonies and 12 (replicated)  
199 lineages: three admixed and three non-admixed lineages on cowpea, and three admixed and  
200 three non-admixed lineages on lentil. This full factorial experimental design allowed us to  
201 compare the evolution and performance of admixed lineages across environments (stressful  
202 versus benign), as well as compare evolution and performance of of admixed versus purebred  
203 lineages within each of those environments. We chose not to conduct the host shift onto  
204 lentil until the F2 generation because F1 hybrids are typically phenotypically uniform and  
205 thus will not reflect the adaptive potential that could emerge after independent assortment  
206 and recombination break down ancestry blocks generating novel genotypic combinations (i.e.  
207 transgressive segregation). All 132 colonies were maintained for at least 20 generations post-  
208 admixture (at least 17 generations post-host shift). After this time colonies were culled via

209 freezing.

## 210 **Population Growth Assays**

211 During the first 400 days after the host shift onto lentil (or until enough beetles emerged  
212 to move the colony into standard culture), we removed all dead adult beetles produced by  
213 each of the 66 lentil colonies. This was done to assess the rate of adaptation to lentil in each  
214 colony, measured by population growth. Every 20 days, beetles from each lentil colony were  
215 separated from the beans using a soil sieve. All live beetles were aspirated from the upper  
216 edge of the sieve and returned to the culture jar to continue laying eggs. This was done to  
217 ensure that population sizes during adaptation to lentil were not altered by our population  
218 growth tracking method. All dead beetles remaining at the bottom of the sieve were removed  
219 and stored at  $-80^{\circ}\text{C}$  until image analysis. Thus, each sample of beetles removed represents  
220 the number of adult beetles that died during the previous 20-day period, and the full set  
221 of such samples for each colony provides an accurate estimate of the cumulative population  
222 size of each colony over time.

223 To assess the number of adult beetles produced by each colony during every 20 days  
224 post host-shift, we used the program **ImageJ** (version 1.52A) (Schneider *et al.*, 2012). Beetle  
225 specimens from each sample were photographed using a Canon EOS M6 camera. Pho-  
226 tographs were first prepared for analysis using the program Adobe Photoshop Elements  
227 2020 Editor to correct uneven lighting and ensure the background color was uniform across  
228 the entire image. This was necessary to ensure that **ImageJ** could accurately differenti-  
229 ate between the color of beetles versus the background sheet. We then used the **analyze**  
230 **particles** function in **ImageJ** to count the number of beetles in each image. The result  
231 of this analysis was a count of the total number of beetles that died during every 20-day  
232 period in each colony post host-shift. As we collected every dead beetle produced by each  
233 jar during each 20 day interval between 60 and 400 days post host shift (unless the colony  
234 was moved into standard culture prior to 400 days post host shift), these population counts

235 represent a complete count of the total number of beetles produced by each colony during  
 236 early adaptation to lentil.

237 We analyzed population growth in both admixed and non-admixed lines using a  
 238 Bayesian generalized linear model. Cumulative count data were assumed to follow a normal  
 239 distribution with  $\mu = \mu^{count}$  and  $\sigma = \sigma^{count}$ . Mean cumulative population count ( $\mu^{count}$ ) was  
 240 assumed to follow a second order polynomial relationship with respect to the number of days  
 241 post host shift such that for non-admixed lineages:

$$\mu^{count} = (\beta_1^{pop} + \alpha_1^{rep})days + (\beta_2^{pop} + \alpha_2^{rep})days^2$$

242 where  $\beta_1^{pop}$  and  $\beta_2^{pop}$  are the effects of time (calculated as the standardized but not centered  
 243 number of days post host-shift) on the mean cumulative number of beetles that emerged for  
 244 each non-admixed population, *pop* is the particular non-admixed population being considered  
 245 (BF, BZ, or CA), *days* is the number of days post host-shift, and  $\alpha_1^{rep}$  and  $\alpha_2^{rep}$  are random  
 246 effects of replicate for each  $\beta$ -term (data from replicates 2 through 10 were used for this  
 247 analysis). Replicate effects were transformed with a sum-to-zero constraint to ensure all  
 248 parameters in the model were identifiable. For admixed lineages,  $\mu^{count}$  was assumed to  
 249 follow the same polynomial relationship shown above except that each slope ( $\beta_1$  and  $\beta_2$ ) for  
 250 admixed populations was assumed to equal the average slope from each parental lineage plus  
 251 an additional effect of admixture, such that:

$$\mu_{count} = \left( \frac{\beta_1^{P1} + \beta_1^{P2}}{2} + \beta_1^{AE} + \alpha_1^{rep} \right) days + \left( \frac{\beta_2^{P1} + \beta_2^{P2}}{2} + \beta_2^{AE} + \alpha_2^{rep} \right) days^2$$

252 where  $\beta^{P1}$  terms are the effects of time on cumulative beetles emerged in the first parental  
 253 lineage,  $\beta^{P2}$  terms are the effects of time on cumulative beetles emerged in the second  
 254 parental lineage, and  $\beta^{AE}$  terms are the additional effects of admixture on the cumulative

255 beetles emerged. Thus, our model included six  $\beta_1$  and six  $\beta_2$  parameters (one slope parameter  
 256 for each of the three parental lineages, and one admixture effect parameter for each of the  
 257 three admixed lineages). Both the  $\beta_1$  and  $\beta_2$  parameters were assigned a normal prior with  $\mu$   
 258  $= 0$  and  $\sigma = 100$ . Raw (not sum-to-zero transformed) random replicate effects (i.e.  $\alpha_1$  and  $\alpha_2$   
 259 parameters) were assigned normal priors with  $\mu = 0$  and  $\sigma = \sigma_1^\alpha$  and  $\sigma_2^\alpha$  respectively. Finally,  
 260 all three sigma parameters ( $\sigma^{count}$ ,  $\sigma_1^\alpha$ , and  $\sigma_2^\alpha$ ) were assigned gamma priors with parameters  
 261  $k = 0.1$  and  $\theta = 0.01$ . This model was written in the language Stan (Stan Development Team,  
 262 2022b) and implemented with the R-interface `rstan` version 2.21.5 (Stan Development Team,  
 263 2022a). We ran 5 chains with a burn-in period of 1,500 steps and 3,000 Hamiltonian Monte  
 264 Carlo (HMC) sampling steps.

## 265 DNA Sequencing, Alignment, and Variant Calling

266 We extracted DNA from between 19-20 beetle specimens each from 78 unique lineage, repli-  
 267 cate, host and generation combinations, for a total of 1536 individuals (Fig. 1). As cowpea  
 268 seed beetles have an XY sex chromosome system and the Y-chromosome is significantly  
 269 reduced in size (Angus *et al.*, 2011; Arnqvist *et al.*, 2023), we chose to sequence only fe-  
 270 male beetles to achieve better coverage of the X-chromosome. We sequenced DNA from  
 271 three time points during our experiment: generation 1 (F1; pre-adaptation), generation 7  
 272 (F7; early adaptation), and generation 20 (F20; late adaptation). From the F1 generation,  
 273 we sequenced only purebred parental cowpea lineages (BF, BZ, and CA) from replicate 1.  
 274 Because our admixed lineages were true F1 hybrids of our parental cowpea lineages, the  
 275 initial allele frequencies of our first generation hybrid lines could be inferred from the allele  
 276 frequencies of these original parental lines. From the early adaptation (F7) generation we  
 277 sequenced replicates 1 to 5 for all cowpea- and lentil-adapted admixed lineages (BF×BZ,  
 278 BF×CA, and BZ×CA) for a total of 30-F7 experimental groups. From our late adaptation  
 279 time point (F20) we again sequenced beetles from replicates 1 to 5 from all admixed lineages  
 280 (both cowpea- and lentil-adapted), as well as all purebred lentil-adapted populations for a

281 total of 45-F20 experimental groups. This sampling scheme allowed us to assess evolution  
282 in purebred lines during adaptation to lentil, evolution in admixed lineages during early  
283 and late adaptation to lentil, as well as evolution during early and late generation admixed  
284 lineages not exposed to a novel host.

285 To extract DNA from beetle specimens, we used Qiagen DNeasy 96 Blood and Tis-  
286 sue Kits. To minimize cross-contamination of DNA, all beetle specimens were washed prior  
287 to DNA extraction. Reduced-representation restriction-fragment-based DNA libraries were  
288 then prepared from extracted DNA using the genotyping-by-sequencing (GBS) library prepa-  
289 ration protocol described in Parchman *et al.* (2012) and Gompert *et al.* (2012) with mod-  
290 ifications from Gompert *et al.* (2014). Briefly, whole-genome DNA was first digested with  
291 MseI and EcoRI enzymes, then ligated to custom barcode sequences and amplified via  
292 PCR. Barcoded and amplified DNA fragments were pooled, purified, and size-selected on  
293 a BluePippin. We selected DNA fragments between 250-350 bp for sequencing. Our DNA  
294 fragment libraries (four libraries total, each with 384 individuals) were sequenced on an Il-  
295 lumina NovaSeq (one full run per library with SP 100 cycles) by the Genomics Core at the  
296 University of Colorado Anschutz Medical Campus.

297 Sequencing resulted in a total of 4,381,945,291 individual reads. We first filtered each  
298 of the Fastq files to remove PhiX sequences. After the removal of PhiX reads, we were left  
299 with a total of 3,539,264,296 reads for alignment. Barcode sequences were then removed  
300 from the remaining reads using a custom perl script, and each read was tagged with the  
301 ID of the beetle from which it came. We aligned DNA reads from our experiment to the  
302 *Callosobruchus maculatus* reference genome (NCBI accession number CASHZR020000000)  
303 using the `bwa aln` algorithm (Li & Durbin, 2009). For this, we set the maximum number of  
304 mismatches allowed per sequence (`-n`) to 5, the seed length (`-l`) to 20, and the maximum  
305 mismatches allowed in the seed sequence (`-k`) to 2.

306 To identify sites with single nucleotide polymorphisms (SNPs), we conducted variant  
307 calling using `bcftools` version 1.16 (Li *et al.*, 2009). We used the original consensus caller

308 (-c) and called only variants for which the posterior probability of the SNP being invariant  
309 was less than 0.01 (-p = 0.01). Variable sites were filtered for quality using custom perl  
310 scripts. In particular, we retained only variable sites with a phred-scale mapping quality  
311 greater than 30, a coverage level equal to or greater than 3072 reads (2× the number of  
312 individuals we sequenced), a minimum of 10 reads for the alternative allele (to filter out  
313 possible errors in sequencing), and representing 80% or more of the individuals we sequenced.  
314 Variable sites with base-quality rank-sum, mapping-quality rank-sum, or read-position rank-  
315 sum test *P*-values less than 0.001, 0.0001, and 0.001 respectively were not retained. After this  
316 initial filtering step, SNPs with a read depth exceeding 48,000, that is, 3 standard deviations  
317 greater than the mean read depth across loci, were also removed. This was done to remove  
318 possible paralogs and gene families from our filtered SNP set. Variable sites located less  
319 than 2 bps apart were also removed. After quality filtering, we were left with 79,079 SNPs  
320 for downstream analysis.

## 321 Population Genetic Analyses

322 In order to obtain robust genotype estimates and quantify patterns of admixture and global  
323 ancestry (i.e. genome-average ancestry), we used the program **entropy** (version 2.0) (Gom-  
324 pert *et al.*, 2014; Shastry *et al.*, 2021). This program is comparable to the admixture model  
325 in **structure**, but with the additional feature of accounting for uncertainty in genotypes,  
326 which are estimated from genotype likelihoods as part of the analysis. Because our experi-  
327 ment used three known parental populations for the production of admixed lineages, we ran  
328 **entropy** only for  $K = 3$  source populations. We ran 20 Markov chain Monte Carlo (MCMC)  
329 chains with 2000 burn-in steps and 2500 sampling steps each, a Dirichlet initialization value  
330 of 50, and a thinning interval of 5. Ancestry proportion estimates generated by **entropy** were  
331 used to determine the degree to which ancestry proportions shifted over time (for example, if  
332 ancestry from one parental lineage was selected against due to incompatibilities or ecological  
333 selection). We also visualized patterns of genetic structure among our parental and admixed

334 lineages by conducting a PCA of the Bayesian genotype estimates from **entropy**. This PCA  
335 was performed from centered but unscaled genotype estimates using R version 4.2.2 (R Core  
336 Team, 2022).

337 We then used the program **popanc** (version 0.1) to estimate population-level, local  
338 ancestry frequencies along chromosomes for each line (Gompert, 2016). This was done to  
339 visualize differences in the frequency of ancestry blocks across the genome and among treat-  
340 ment groups. This program uses a continuous correlated beta process model for inferring  
341 ancestry, and is particularly well-suited for inferring ancestry in hybrid populations that do  
342 not experience ongoing gene flow with parental populations and for which genome stabi-  
343 lization is not yet complete (Gompert & Buerkle, 2013; Gompert, 2016). We ran **popanc**  
344 using the genotype estimates from **entropy**. We only included SNPs assigned to one of the  
345 ten *C. maculatus* chromosomes (9 autosomes and the X chromosome; 72,583 of the 79,079  
346 SNPs) and for which the absolute difference in initial allele frequencies between parental  
347 lineages (BF versus BZ, BF versus CA, or BZ versus CA) was greater than 0.2. This was  
348 done to ensure that only the loci that were informative of population ancestry were used for  
349 local ancestry analysis. We chose to have **popanc** estimate the scale parameter for the beta  
350 process model (-s) and set a uniform prior on this parameter, U(lower = 1, upper = 100,000)  
351 (here measured in bps). We set a maximum locus distance (-d) to one Megabase and the  
352 maximum number of SNPs per locus (-n) to 15. We ran two MCMC chains for each admixed  
353 lineage with each comprising a 10,000 step burn-in and 30,000 sampling steps with a thinning  
354 interval (-t) of 5.

355 To estimate allele frequencies within each unique lineage, replicate, host and gener-  
356 ation combination, we used the program **estpEM** (Soria-Carrasco *et al.*, 2014). **estpEM** uses  
357 the expectation-maximization algorithm described in Li (2011) and accounts for uncertainty  
358 in genotypes during allele frequency estimation. For this analysis, we used a convergence  
359 tolerance of 0.001 and allowed for a maximum of 20 iterations. F1 allele frequency estimates  
360 from **estpEM** were used to calculate Nei's  $F_{ST}$  between pairs of non-admixed, parental lin-



361 eages (i.e. BF, BZ, and CA cowpea lineages) in order to determine the degree of genetic  
 362 differentiation among our parental lineages. We then computed allele frequency change for  
 363 all of our sequenced experimental groups. Allele frequency change, or  $\Delta p$ , was calculated  
 364 as  $\Delta p = p_t - p_0$  where  $p_t$  is the frequency of an allele at time point  $t$  (7 or 20) and  $p_0$  is  
 365 the initial frequencies of allele. Because we formed true F1 hybrids to establish our admixed  
 366 populations, the initial allele frequencies of our admixed lineages should be a simple average  
 367 of the allele frequencies of the parent populations. As such, we estimated the allele frequen-  
 368 cies of our F1 hybrid lineages by taking the average of the allele frequencies of their parental  
 369 lineages.

370 Finally, we used `varne` to estimate the contemporary, variance effective population  
 371 size ( $N_e$ ) of each colony based on patterns of allele frequency change during the experiment  
 372 (Jorde & Ryman, 2007; Gompert & Messina, 2016; Rêgo *et al.*, 2019). This allowed us to  
 373 (i) estimate the severity of population bottlenecks experienced by both admixed and non-  
 374 admixed lines during adaptation and (ii) obtain estimates of  $N_e$  to parameterize the null  
 375 model of expected evolutionary change by genetic drift described in the next section. We  
 376 estimated variance  $N_e$  between generation F1 and generation F20, and conducted all `varne`  
 377 analyses with an approximate census size (`-n`) of 2000 beetles and 1000 Bayesian bootstrap  
 378 replicates (`-x`).

## 379 **Testing for Repeated Adaptive Evolution**

380 We constructed a null model to determine whether the observed degree of allele frequency  
 381 change for each locus in each line was greater than expected by genetic drift alone. We  
 382 modeled evolution by drift using a beta-distribution approximation to a Wright-Fisher model  
 383 (Ewens & Ewens, 2004; Gaggiotti & Foll, 2010; Rêgo *et al.*, 2019). Here, the the probability  
 384 of allele frequency  $p_t$  at time  $t$  follows a beta distribution with  $\alpha = p_0 * (1 - F)/F$  and  $\beta =$   
 385  $(1 - p_0) * (1 - F)/F$ , where  $F = 1 - (1 - \frac{1}{2N_e})^t$ . Thus, the magnitude of change by drift depends  
 386 on effective population size, time, and initial allele frequency. We parameterized the null

387 model using the actual number of elapsed generations ( $t$ ), our estimates of variance effective  
 388 population size ( $N_e$ ) from `varne`, and our maximum likelihood estimates of allele frequencies  
 389 ( $p_0$ ). We constrained allele frequencies to be between 0.01 and 0.99 for numerical stability;  
 390 greater precision than this would also be difficult to justify from our sample sizes. We  
 391 converted the one-tailed probability from the beta probability distribution function (`pbeta`  
 392 in R) to a two-tailed  $P$ -value by taking  $\min[P_{\text{beta}} * 2, (1 - P_{\text{beta}}) * 2]$ . We interpret these  
 393  $P$ -values as measures of the evidence against the null hypothesis that evolution occurred  
 394 only by drift, and thus as evidence that evolution was directly or (more likely) indirectly  
 395 (via linkage disequilibrium) effected by selection.

396 We conducted a series of analyses to quantify the extent to which the same loci  
 397 exhibited the greatest evidence of non-neutral evolution (as captured by our null-model  $P$ -  
 398 values) (i) across replicate lines of the same lineage and treatment and (ii) between different  
 399 pairs of lineages or treatments. Thus, we were interested in both whether evolutionary change  
 400 during adaptation was more repeatable under some conditions than others and whether  
 401 evolution was more repeatable for certain pairs of conditions (e.g., admixed and non-admixed  
 402 lines adapting to the same host or admixed lines adapting to different hosts) than others. We  
 403 first used the program `picmin` to test for repeated, non-neutral evolution at the SNP level  
 404 for each treatment and sequenced time point (Booker *et al.*, 2023). This approach works  
 405 by identifying loci (here SNPs) that consistently fall in the tails of an empirical  $P$ -value  
 406 distribution across a set of populations or species. Such patterns are indicative of a repeated  
 407 association with adaptation. Most applications of this method have considered somewhat  
 408 distantly related taxa and have taken gene or window-based approaches (e.g., Nocchi *et al.*,  
 409 2024; Whiting *et al.*, 2024). Here, we apply this to replicates of the same source population  
 410 and focus on individual SNPs (which are necessarily shared across the set of replicates).

411 For these analyses, we converted the null-model  $P$ -values from each line to an em-  
 412 pirical (ranked order)  $P$ -value distribution. We excluded SNPs with initial minor allele  
 413 frequencies less than 0.01 from this analysis as uncertainty in the precise allele frequencies

414 for such rare alleles could have a disproportionate affect on the evidence against neutral  
415 evolution. We generated the null correlation matrix for each set of replicate lines using the  
416 `GenerateNullData` function with 10,000 replicate draws,  $a = 0.3$ , and  $b = 0.5$ . We then  
417 applied the `picmin` function to the empirical  $P$ -values each set of replicate lines (i.e. the five  
418 lines for each combination of lineage, generation and host). We applied a false discovery rate  
419 (FDR) correction to the `picmin`  $P$ -values to designate SNPs with significant evidence of re-  
420 peated non-neutral evolution for each treatment group (i.e.  $P < 0.05$  after FDR correction)  
421 (Benjamini & Hochberg, 1995).

422 We then asked whether the same SNPs showed evidence of repeated non-neutral evo-  
423 lution in different pairs of treatment groups (combinations of lineage, generation and host).  
424 For this, we identified the 5% of SNPs with the lowest  $P$ -values from `picmin`, regardless of  
425 whether these were less than 0.05 after FDR correction (our 5% cutoff is approximately of  
426 the same order as the average number of SNPs with significant evidence of repeated evolu-  
427 tion in each treatment). Next, for each pair of treatment groups, we computed the overlap  
428 between these sets of SNPs, that is, the number of SNPs in the top 5% for pairs of treatment  
429 groups. This served as our observed measure of repeated-evolution SNP-sharing between  
430 treatments and specifically captured the extent to which SNPs repeatably showing evidence  
431 of non-neutral evolution within treatment groups were shared between treatment groups. We  
432 generated null expectations for the overlap expected by chance by repeatedly randomizing  
433 `picmin`  $P$ -values among SNPs; the randomization procedure preserved information on which  
434 SNPs were included in the `picmin` analysis for each treatment group. This was done 1000  
435 times for each treatment group comparison. We used this null distribution to calculate a  
436 randomization test  $P$ -value for whether the observed overlap exceeded chance expectations  
437 and to calculate the X-fold enrichment of the observed value relative to the null, that is, the  
438 ratio of the observed overlap to null expectations, which serves as a quantitative measure of  
439 the extent of repeatability between treatments. These analyses were done in R version 4.4  
440 (R Core Team, 2022).

## 441 Results

### 442 Population Growth and Adaptation to Lentil

443 Despite poor initial survival, 64 out of our 66 experimental lentil lines successfully adapted  
444 to this novel, stressful host. The two lines that did not adapt to lentil were BF replicates  
445 6 and 7. The BF lineage showed the slowest cumulative growth rate, while the BF×CA  
446 admixed lineage showed the highest (Fig. 2 and Table S1). We also saw a strong incubator  
447 effect, with lines 2 through 5, which were housed in the first incubator, experiencing more  
448 rapid cumulative growth than populations 6 through 11, which were housed in the second  
449 incubator (Fig. 2). Overall, we found that admixture facilitated adaptation to lentil, with  
450 higher cumulative growth rates occurring in admixed than non-admixed lineages (Fig. 2 and  
451 Table S1).

452 Results from our Bayesian second-degree polynomial model for cumulative population  
453 growth showed a strong signal for evolutionary rescue in all our lentil-adapted lineages.  
454 Values of  $\beta_1$  indicate the slope of the cumulative growth curve at time  $t = 0$ . Thus,  $\beta_1$  values  
455 can be interpreted as an estimate of the average reproductive rate of each lineage at time  $t =$   
456 zero. The higher the value of  $\beta_1$ , the higher the initial reproductive rate on lentil. The 95%  
457 credible intervals for  $\beta_1$  for the BF, BZ, and BF×BZ population all overlapped zero (Table  
458 S1). This suggests that at time  $t = 0$ , the reproductive rate (as measured by the average  
459 number of adult offspring produced per day) in these populations was not high enough to  
460 ensure population persistence on lentil.

461 Alternatively,  $\beta_1$  values could also be interpreted as a measure of how long it would  
462 take for a given founding population of parent beetles to produce enough offspring to fully  
463 replace itself, assuming a parental death rate of zero and non-overlapping generations (i.e.  
464 generation time). For the CA, BF×CA, and BZ×CA lineages,  $\beta_1$  ranged from 4.7 and 8.6,  
465 suggesting that the average reproductive rate at time  $t = 0$  in these lineages was between  
466 4.7 and 8.6 adult offspring per day (see Table S1). At this reproductive rate, it would

467 hypothetically take between 200 to 500 days for our founding populations of 2000 adult  
468 beetles to produce 2000 adult offspring. This, of course, would not be possible in reality  
469 as adult seed beetles have limited adult lifespans (less than 10 days) and the majority of  
470 first-generation offspring surviving on lentil expected to emerge within 100 days. Thus,  
471 even for the three lineages with  $\beta_1$  values credibly greater than zero, the initial reproductive  
472 rate estimated by our model was not high enough to suggest that these populations would  
473 produce enough offspring to prevent an initial population decline. Our model results for  $\beta_1$   
474 indicate that, on average, all three admixed and all three non-admixed lineages are expected  
475 to undergo an initial demographic decline, consistent with the first stage of evolutionary  
476 rescue.

477 The second slope parameter from our Bayesian model,  $\beta_2$ , is a measure of growth  
478 rate. A  $\beta_2$  value of zero indicates that population size will remain constant with respect  
479 to time (in other words, the population size is stable and no growth occurs), while any  
480 value of  $\beta_2$  greater than zero indicates exponential growth, meaning population size will  
481 increase with time. A negative value of  $\beta_2$ , meanwhile, indicates that population size will  
482 decrease with time. Values of  $\beta_2$  in admixed populations were calculated as the average of  $\beta_2$   
483 values for each parental lineage plus an effect of admixture ( $\beta_2^{AE}$ ).  $\beta_2^{AE}$  values of zero would  
484 indicate that the cumulative growth rate in admixed populations was simply the mean of  
485 the parental populations' cumulative growth rates. In other words, a  $\beta_2^{AE}$  of zero indicates  
486 that the cumulative growth rate of admixed populations falls directly between those of its  
487 parents. The effect of admixture for  $\beta_2$  in our linear model was credibly greater than zero  
488 for all admixed lineages (Table S1), suggesting that growth rates in all three of our admixed  
489 lineages were greater than the average of their parents' growth rates. Notably, values  
490 of  $\beta_2$  were credibly greater than zero for all populations, both admixed and non-admixed,  
491 indicating that all six populations on average were expected to rebound from their initial  
492 demographic decline on lentil.

## Population Structure and Evolutionary Change

Our source cowpea-adapted beetle lineages from Burkina Faso (BF), Brazil (BZ), and California (CA) showed a moderate to strong degree of genetic differentiation from one another. The degree of genetic differentiation between our African population and our two American populations ( $F_{ST}$  for BF and BZ and for BF and CA = 0.20) was twice as high as the degree of differentiation between the two American populations ( $F_{ST}$  for BZ versus CA = 0.09). This result was recapitulated in a PCA (Fig. 3a), with PC1 separating the BF lineage from the BZ and CA lineages, and PC2 separating the BZ and CA lineages. As expected, our three admixed lineages (BF×BZ, BF×CA, and BZ×CA) clustered directly between their two parental populations. One of our lentil-adapted BF lines, BF replicate 5 generation 20, clustered with the BF×BZ admixed lines rather than with the BF purebred lines, indicating that this BF replicate was likely contaminated with BZ beetles at some point during the experiment and underwent admixture. As such, this single BF replicate was removed from all downstream analyses. We observed possible, weaker evidence for contamination of lentil-adapted BF×BZ replicate 2 with CA, but the similarity between BZ and CA makes this less clear and we thus chose to retain this replicate.

Global ancestry estimates also showed clear evidence of admixture consistent with expectations based on their hybrid ancestries (Fig. 3b). Comparison between F7 and F20 generation lentil-adapted hybrids from the BF×BZ and BF×CA lineages showed that global BF ancestry declined over time. While F1 hybrids would have received exactly 50% of their genome from each parental lineage, the mean BF ancestry in F20 admixed lentil lines ranged between 38-45%, a 5-12 percentage point decline in BF ancestry over the course of adaptation to lentil. The F20 admixed cowpea lines, in contrast, showed mean global BF ancestry values between 52-55%. This indicates possible selection against BF ancestry during adaptation to lentil.

Estimates of local ancestry–ancestry block frequencies along chromosomes–also re-

519 vealed a decline in BF ancestry in lentil-adapted admixed lineages, but not in lineages on  
 520 cowpea (Figs. 4 and S1–S3). Reduced BF ancestry was especially evident on chromosome 9,  
 521 and this was especially true for the BF×CA lines (Figs. 4 and S2). In contrast, local ancestry  
 522 frequencies in lentil and cowpea-adapted BZ×CA lines were ~0.5 across most of the genome.  
 523 With the exception of lentil-adapted BF×BZ, patterns of local ancestry were similar among  
 524 replicate lines. For lentil-adapted BF×BZ, BF ancestry was low on most chromosomes in  
 525 replicates 2 and 3, whereas BF ancestry was only notably reduced on chromosome 9 in  
 526 replicates 1, 4 and 5 (Fig. S2 and S3).

527 In addition to the changes in ancestry in admixed populations described above, we  
 528 documented pervasive, genome-wide evolutionary changes in all populations over the course  
 529 of this experiment. Mean allele frequency changes (across SNPs and replicates) ranged from  
 530 0.046 to 0.072 by generation F7 and 0.051 to 0.094 by generation F20 (Figs. 5, 6 and S4-  
 531 S6). The biggest changes occurred in lentil-adapted BF, and allele frequency changes were  
 532 generally larger in lentil-adapted than cowpea-adapted lines. Patterns of allele frequency  
 533 change varied across the genome. For example, we detected peaks of more pronounced  
 534 change on chromosome 1 in lentil-adapted BZ, CA, BF×BZ and BZ×CA (Figs. 5 and 6).  
 535 Similarly, peaks of pronounced allele frequency change were visible on much of chromosome  
 536 9 in the lentil-adapted BF×BZ and BZ×CA lineages (Figs. 6 and S5).

537 Variance effective population sizes ( $N_e$ ) estimated from the F1 to F20 generations  
 538 varied from a minimum of 38.7 (95% credible interval [CI] 38.0-39.3) in BF replicate 4  
 539 on lentil to a maximum of 222.4 (95% CI 216.3-229.2) in BZ×CA replicate 5 on cowpea,  
 540 consistent with the documented degree of genome-wide allele frequency change (Table S2).  
 541 All  $N_e$  estimates were considerably lower than the founding population size of the colonies in  
 542 our experiment (1000 beetles per colony). These estimates varied considerably both across  
 543 hosts and source populations, but were generally higher for cowpea lines than lentil lines.

## 544 **Patterns of Repeated Adaptive Evolution**

545 We found genome-wide evidence of allele frequency change beyond that predicted by the  
546 null Wright-Fisher models, with the most pronounced evidence of exceptional evolutionary  
547 change often on chromosomes 1 and 9 (Figs. 7, 8, and S7-S9). Wide peaks of non-neutral  
548 evolution were especially evident on chromosome 1 in lentil-adapted BZ, CA and BZ×CA  
549 and on chromosome 9 in lentil-adapted BF×BZ and BF×CA (Figs. 7 and 8).

550 In some cases, SNPs showed strong evidence of selection in only a subset of replicate  
551 lines (see, for example, the large peak on the right side of chromosome 1 for lentil-adapted  
552 BF×BZ, Fig. 8a). Nonetheless, the `picmin` analyses identified hundreds to thousands of  
553 SNPs associated with repeated adaptive evolution in each of the treatment groups (Fig.  
554 9a). In general, more SNPs showed significant evidence of repeated, adaptive evolution  
555 in lentil-adapted lines than cowpea-adapted lines. The effect of admixture was less clear.  
556 Repeatability was highest in the BF×BZ and BF×CA lineages, followed by non-admixed  
557 BZ and CA, and then admixed BZ×CA and BF (Fig. 9a). Thus, repeatability was high  
558 in admixed lineages that included BF as one of the source lineages but especially low in  
559 non-admixed BF (with the caveat that the latter is partially explained by having four rather  
560 than five replicates). Finally, in the admixed lineages, repeatability was higher in the F7  
561 generation than in the F20 generation. Fewer than 5% of SNPs exhibited significant ( $P <$   
562  $0.05$  after FDR correction) evidence of repeated adaptive evolution for most chromosomes  
563 and treatment groups, but repeatability was higher for some chromosomes. For example,  
564 repeatability was often accentuated in lentil-adapted lines on chromosomes 1 or 9, and in  
565 some cases this was even true for cowpea-adapted lines (e.g., chromosome 9 for BF×CA on  
566 cowpea) (see Fig. 9b). Interestingly, CA was unique in having a especially high proportion  
567 of repeated adaptation SNPs on chromosome 5.

568 Most pairs of treatment groups and time periods (97 out of 105) showed more evidence  
569 of shared, repeated evolution beyond that expected by chance (Table S3, S4 and Fig. 10).



570 Overall, the highest excess of shared, repeated evolution was found for comparisons involving  
571 BF×BZ and BF×CA. For these comparisons, 5.89 to 11.60 times more SNPs than expected  
572 were among the top 5% repeated evolution in both treatment groups (or time periods). As  
573 expected, the highest overlap was for subsequent time points within the same treatment  
574 group (Table S3, S4 and Fig. 10). The effects of admixed versus non-admixed and same  
575 versus different host on repeated evolution between treatment groups were more nuanced.  
576 Excess overlap was higher between lentil-adapted BF×BZ and BF×CA (8.20 to 10.48×)  
577 than between either (ii) cowpea-adapted BF×BZ and BF×CA (7.68 to 8.79×) or (iii) lentil  
578 and cowpea-adapted groups from either BF×BZ or BF×CA (6.20 to 8.97×) (Table S3).  
579 In contrast, evidence of shared repeated evolution SNPs was weaker between BZ×CA and  
580 BF×BZ or BF×CA (1.73–3.81×). We detected notable parallelism between non-admixed  
581 BF on lentil and all admixed lineages involving BF (BF×BZ and BF×CA), especially on  
582 lentil (5.55 to 5.91×) (Table S3, S4 and Fig. 10), whereas lentil-adapted BZ×CA exhibited  
583 greater parallelism with non-admixed BZ and CA on lentil (3.85 to 5.84×).

## 584 Discussion

585 In this experiment, we assessed patterns of repeated evolution across admixed versus non-  
586 admixed seed beetles during adaptation to a novel, stressful host: lentil. We found that  
587 admixture facilitated adaptation to lentil, with the BF lineage showing the slowest rate of  
588 cumulative population growth during adaptation to this novel host, but evolutionary rescue  
589 occurred in almost all lines, and was thus a repeatable evolutionary outcome. Genomic  
590 analyses revealed that levels of parallelism varied among lineages in a nuanced way, such that  
591 the most SNPs were repeatedly associated with lentil adaptation in the the admixed lineages  
592 BF×BZ and BF×CA (>2000), followed by two non-admixed lineages (BZ and CA; ~1800  
593 SNPs), and then the other admixed and non-admixed lineages (BZ×CA and BF; <1000  
594 SNPs). In other words, repeatability was highest in admixed lineages involving BF and non-

595 admixed lineages excluding BF. SNPs on two chromosomes, 1 and 9, exhibited the highest  
596 average levels of evolutionary change and non-neutral evolution in our experiment. We found  
597 a large spike of allele frequency change on chromosome 1 in many lineages associated with  
598 adaptation to lentil, suggesting that adaptation to this novel host is being driven at least in  
599 part by selection on alleles that are adaptive in both admixed and non-admixed lineages. We  
600 further found evidence for selection against BF ancestry on chromosome 9 across both hybrid  
601 lineages derived from BF parents (BF×BZ and BF×CA), indicating that certain regions  
602 of the Burkina Faso genome are likely globally maladaptive on lentil. This same region of  
603 chromosome 9 in the non-admixed BF lineage showed moderate evidence of exceptional allele  
604 frequency change during adaptation to lentil, again suggesting that certain alleles carried by  
605 the BF lineage are globally maladaptive on lentil, regardless of admixture status. Finally,  
606 we found a moderate degree of parallelism in evolutionary change between admixed lineages  
607 adapted to lentil versus cowpea, suggesting that even under extreme ecological selection, the  
608 purging of hybrid incompatibilities still contributes to the degree of evolutionary parallelism  
609 observed in admixed lineages. We discuss the implications of these results in greater detail  
610 below.

## 611 **Admixture facilitates adaptation to lentil**

612 Interestingly, it appears that the African lineage (Burkina Faso or BF) showed the poorest  
613 capacity to adapt to lentil. The Burkino Faso lineage is from the heart of the purported  
614 ancestral range of cowpea seed beetles (Kébé *et al.*, 2017), and as such might be expected to  
615 harbor greater genetic diversity than American populations, which were transported across  
616 the world via trade and may have undergone significant population bottlenecks during estab-  
617 lishment in new locations. Conversely, cowpea is a crop of particular importance in Africa  
618 and is widely grown (Kpoviessi *et al.*, 2019), meaning cowpea may have been the only host  
619 encountered by the wild Burkina Faso seed beetle population. Cowpea is less widely grown  
620 in the Americas, so it is possible that the two American lineages (Brazil and California) had

621 previous exposure to lentil or to other legume species more commonly grown in these regions,  
622 potentially increasing their ability to adapt to novel hosts (but see Messina & Jones, 2009).  
623 Alternatively, as all of our lineages have been reared in captivity for many generations, it is  
624 possible the Burkina Faso lineage (which was originally collected in 1989; see Messina, 1993)  
625 has simply lost some of its original diversity via genetic drift or adaptation to captivity, and  
626 its poor adaptive capacity on lentil is simply a reflection of this laboratory history.

627         It is also possible that other environmental factors alter the adaptive capacity of  
628 different lineages of seed beetles on lentil. Despite using very similar models of Percival  
629 incubators for this experiment, maintaining the same temperature and day cycle in both, as  
630 well as running a dehumidifier full time in both incubators, we nevertheless saw substantial  
631 incubator effects across our treatment groups. Replicates 1 through 5 were kept in our first  
632 incubator, while replicates 6 through 11 were kept in the second. The first incubator was  
633 prone to periods of higher humidity while the second stayed drier during the course of the  
634 experiment. Adaptation proceeded much more rapidly in the first incubator (see Fig. 2),  
635 and differences in the rate of adaptation across lineages were far less pronounced. Humidity  
636 is strongly affected by the total number of colonies in each incubator due to the amount of  
637 metabolic water produced by larvae (Bhattacharya *et al.*, 2003), and our incubators were  
638 especially prone to humidity spikes during the pupation stage. Humidity has a strong effect  
639 on development time and survival in *C. maculatus* (Mainali *et al.*, 2015; Umoetok Akpas-  
640 sam *et al.*, 2017) with the development being the fastest at humidities between 75-80%.  
641 Higher humidity appears to increase survival on lentil, suggesting that perhaps differences in  
642 adaptive capacity of our parental lineages on lentil could be related to not just the ability to  
643 metabolize the novel host, but also their degree of adaptation to low-humidity environments.  
644 Further work is warranted to determine how these lineages differ in their survival at various  
645 humidity levels, and how the environmental effects of humidity and host interact to affect  
646 survival.

## 647 **Genetics and repeatability of adaptation to lentil**

648 Numerous SNPs spanning much of chromosomes 1 and 9 were repeatedly associated with  
649 adaptation to lentil. These wide and pronounced peaks of association suggest a major role for  
650 linkage disequilibrium and linked (indirect) selection in driving patterns of allele frequency  
651 change during adaptation. Some of this linkage disequilibrium likely resulted from admix-  
652 ture (Falush *et al.*, 2003), but these patterns of change were not only observed in admixed  
653 populations. Thus, these results suggest selection on a few, large regions of reduced recom-  
654 bination, which we hypothesize correspond with large structural variants (i.e. chromosomal  
655 rearrangements). Many other recent studies have documented rapid or repeated adaptive  
656 evolution in involving structural variation, suggesting this might be a general phenomenon  
657 (Todesco *et al.*, 2020; Akopyan *et al.*, 2022; Ma *et al.*, 2024; Nosil *et al.*, 2024; Battlay *et al.*,  
658 2025).

659 Genomic patterns of adaptation to lentil were similar in the BZ, CA and BZ×CA  
660 lineages. The BZ and CA populations are more closely related to each other than either  
661 is to BF ( $F_{ST} \sim 0.1$  versus 0.2). Thus, these documented patterns of repeated evolution at  
662 the genetic level are consistent with the general pattern that gene reuse during adaptation  
663 declines with divergence time or genetic dissimilarity (Conte *et al.*, 2012; Chaturvedi *et al.*,  
664 2022; Bohutínská & Peichel, 2024). Our results also suggest that same alleles can contribute  
665 to lentil adaptation in admixed and non-admixed lineages and thus that the effects of these  
666 alleles do not necessarily depend strongly on genetic background. Likewise, we found some  
667 consistency in adaptation for BF and the admixed lineages BF×BZ and BF×CA, though the  
668 contribution of chromosome 9 was more pronounced in the admixed lineages than purebred  
669 BF. Moreover, the weaker signal on chromosome 1 for BF suggests that the hypothesized  
670 large structural variant on chromosome 1 in BZ and CA might be absent from BF. However,  
671 testing this hypothesis requires additional data and analyses (e.g., whole genome comparative  
672 alignments) or experiments, especially given the large bottleneck (and associated genome-  
673 wide changes) caused by the initial host shift to lentil.

674 Lastly, we found that the degree of parallelism during adaptation to lentil across  
675 replicates in the Brazil (BZ) and California (CA) purebred lineages was higher than the level  
676 of parallelism observed in admixed BZ×CA on lentil. This could simply be a byproduct of  
677 admixture: if transgressive segregation led to a greater variance in genotypes in admixed  
678 populations, then it might be more likely that different genomic backgrounds would survive  
679 the severe population bottleneck imposed by adaptation to lentil in different replicates of  
680 admixed populations. This could lead to a decrease in the predictability of evolution in  
681 admixed lineages during adaptation to extreme environments.

## 682 **Hybrid incompatibilities contribute to parallelism**

683 Despite the strong selective pressure imposed by lentil, the overall level of parallelism be-  
684 tween cowpea- and lentil-adapted lineages was still reasonably high (3.90 to 8.97× higher  
685 than expected by chance). Shared peaks of (non-neutral) allele frequency change between  
686 BF×BZ and BF×CA on both lentil and cowpea lineages suggest that there may be hybrid  
687 incompatibilities associated with the BF lineage that are shared across hybrid types. Taken  
688 together, this evidence suggests that even in the face of strong ecological stress, hybrid  
689 incompatibilities may still play a major role in driving evolutionary change in admixed pop-  
690 ulations. This is consistent with results from other studies on admixed lineages in natural  
691 or neutral environments (Matute *et al.*, 2020; Langdon *et al.*, 2022; Kato *et al.*, 2024; Owens  
692 *et al.*, 2025).

## 693 **Conclusion**

694 In conclusion, we found that admixture facilitated adaptation to lentil, and that adaptation  
695 to lentil in cowpea seed beetles is driven in part by selection on globally-adaptive alleles  
696 in both admixed and non-admixed lineages. We also found evidence that certain regions  
697 of genome from the African lineage (BF) appear to be globally maladaptive on lentil, and

698 this led to parallel selection against BF ancestry in lentil-adapted lineages across hybrid  
699 types. Finally, we saw a moderate degree of parallelism in evolutionary change between  
700 admixed lineages adapted to lentil versus cowpea, suggesting that even during evolutionary  
701 rescue, the purging of hybrid incompatibilities may still be a major contributor to patterns  
702 of evolutionary parallelism observed in admixed lineages.

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## 953 **Data Accessibility and Benefit-Sharing**

### 954 Data Accessibility Statement

955 Raw sequence reads have been deposited in NCBI's SRA (BioProject PRJNA1232334).  
956 Scripts, ecological data and downstream genetic data will be available from Dryad (DOI  
957 pending). Current versions of scripts and ecological data are available on **GitHub** for review  
958 (<https://github.com/zgompert/CmacAdmix>).

### 959 Benefit-Sharing Statement

960 Benefits from this research accrue from the sharing of our data and results on public databases  
961 as described above.

## 962 **Author Contributions**

963 AS and ZG designed the research. AS performed the research. AS, BK and ZG analyzed  
964 the data. AS and ZG wrote the paper.

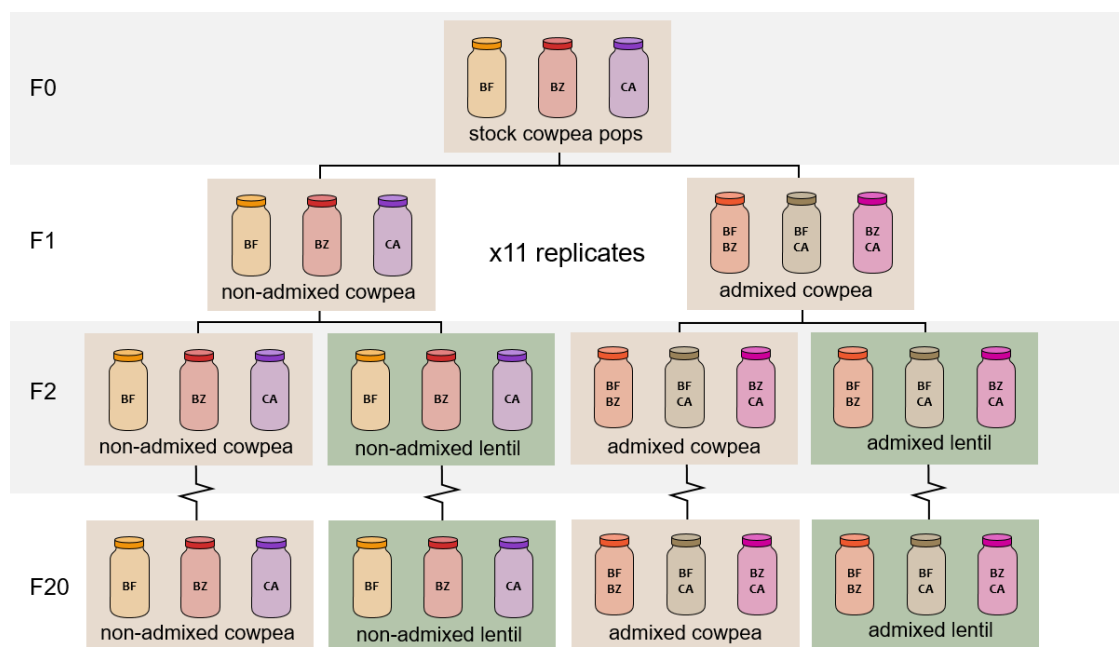
965 **Figures**

Figure 1: Overview of the experimental design. We evolved 11 replicate lines from each of six lineages—Burkina Faso (BF), Brazil (BZ), California (CA), and the admixed lineages BF×BZ, BF×CA and BZ×CA—on an ancestral host, cowpea, and a novel, stressful host, lentil for 20 generations. We generated DNA sequence data from five replicate lines from each experimental group (lineage and host). Samples were sequenced from the F20 (all lineages) and F7 (only the admixed lineages) generations, along with a single replicate of each of the stock source lineages (BF, BZ and CA).

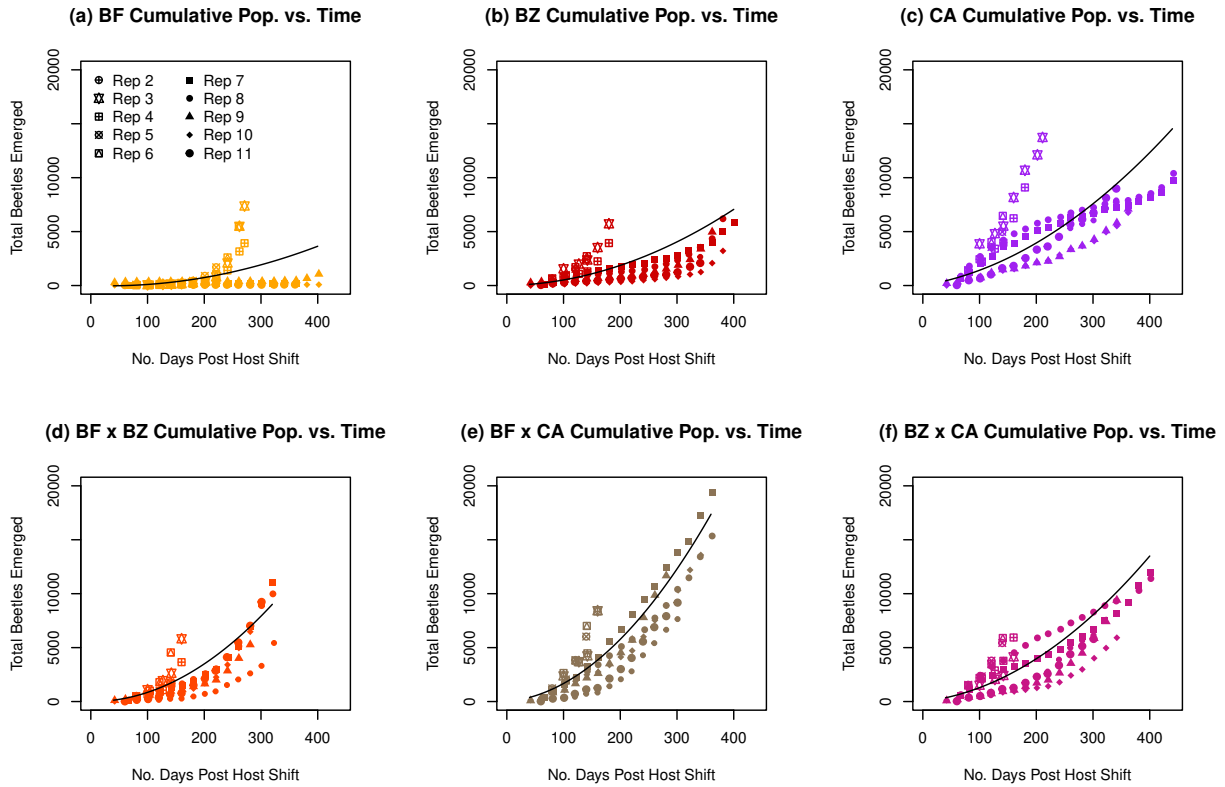


Figure 2: The cumulative number of beetles that emerged from each lentil colony over time by lineage and replicate. Non-admixed lineages are shown in panels (a) through (c), and admixed lineages are shown in panels (d) through (f). Each individual data point represents the total number of beetles produced by a given colony between time  $t = 0$  and time  $t$ , not the population size at time  $t$ . In other words, our plots represent cumulative growth, or the sum of population growth. Thus, a linear relationship between cumulative growth and time would represent a population whose size remains constant with respect to time, while a concave up curve represents population growth over time, and a concave down curve represents a population that is decreasing in size with time. Data points from each individual replicate are represented by point shape. Replicates two through five (hollow point shapes) were all maintained in one incubator, while replicates six through eleven were maintained in a second incubator at the same temperature and day cycle. The average cumulative growth for each lineage fit by our Bayesian model are shown as black curves on each panel.

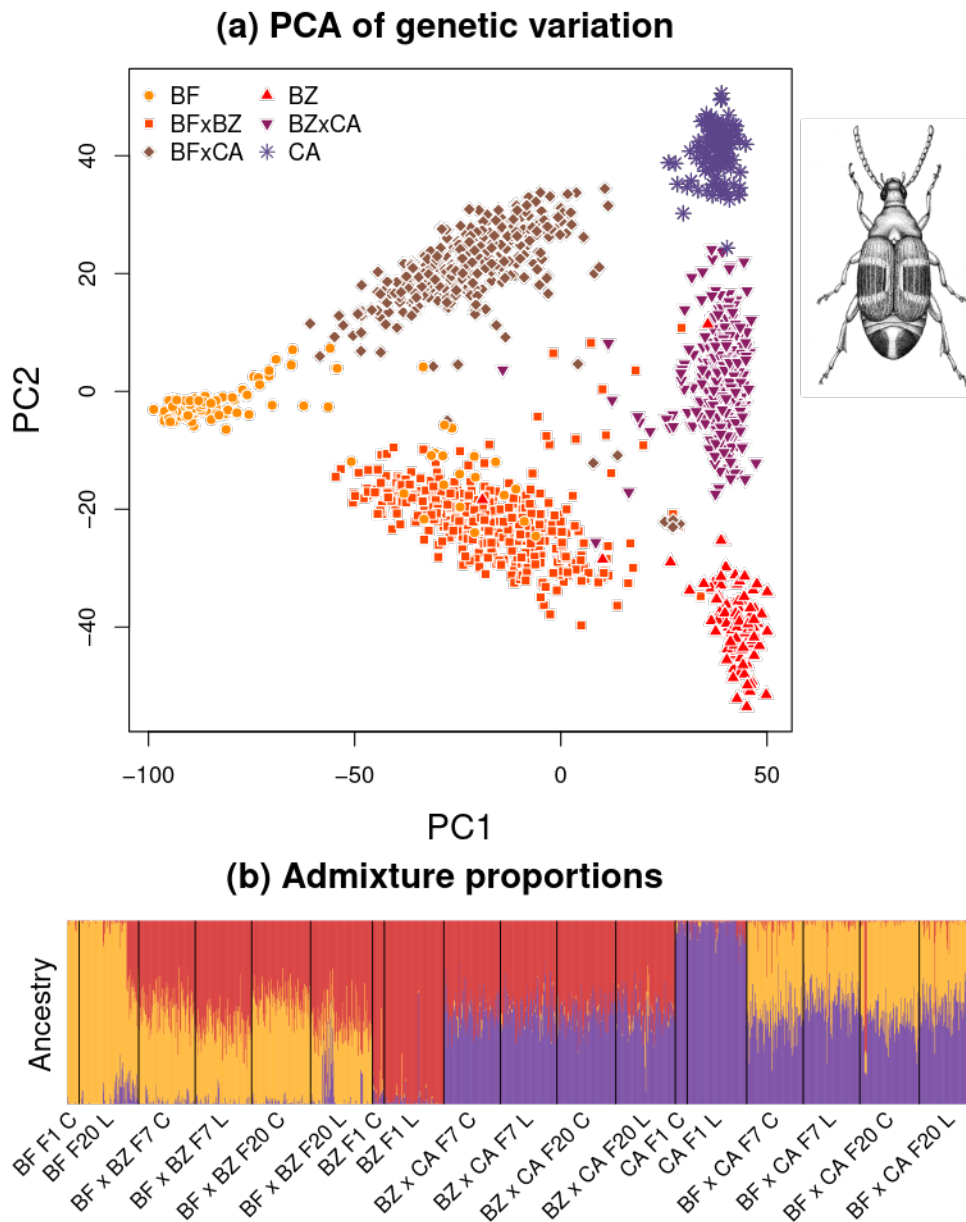


Figure 3: Genetic variation in the experimental *C. maculatus* populations. (a) Principal component analysis of (unscaled) genotype estimates. Each point represents one of the 1536 beetles we sequenced in this study. This includes F7 and F20 generation admixed beetles adapted to both lentil and cowpea from replicates 1 through 5, as well as F1 generation non-admixed beetles adapted to cowpea and F20 generation non-admixed beetles adapted to lentil from replicates 1 through 5. Each unique lineage (BF, BZ, CA, BF×BZ, BF×CA, and BZ×CA) is represented by a unique color×shape combination on the PCA. (b) Admixture proportions for each individual estimated with *entropy*. Vertical bars represents global ancestry proportions for each of the beetles sequenced. Burkina Faso (BF) ancestry is shown in light orange, Brazil (BZ) ancestry in red, and California (CA) ancestry in purple; C and L denote cowpea-adapted and lentil-adapted lineages, respectively.

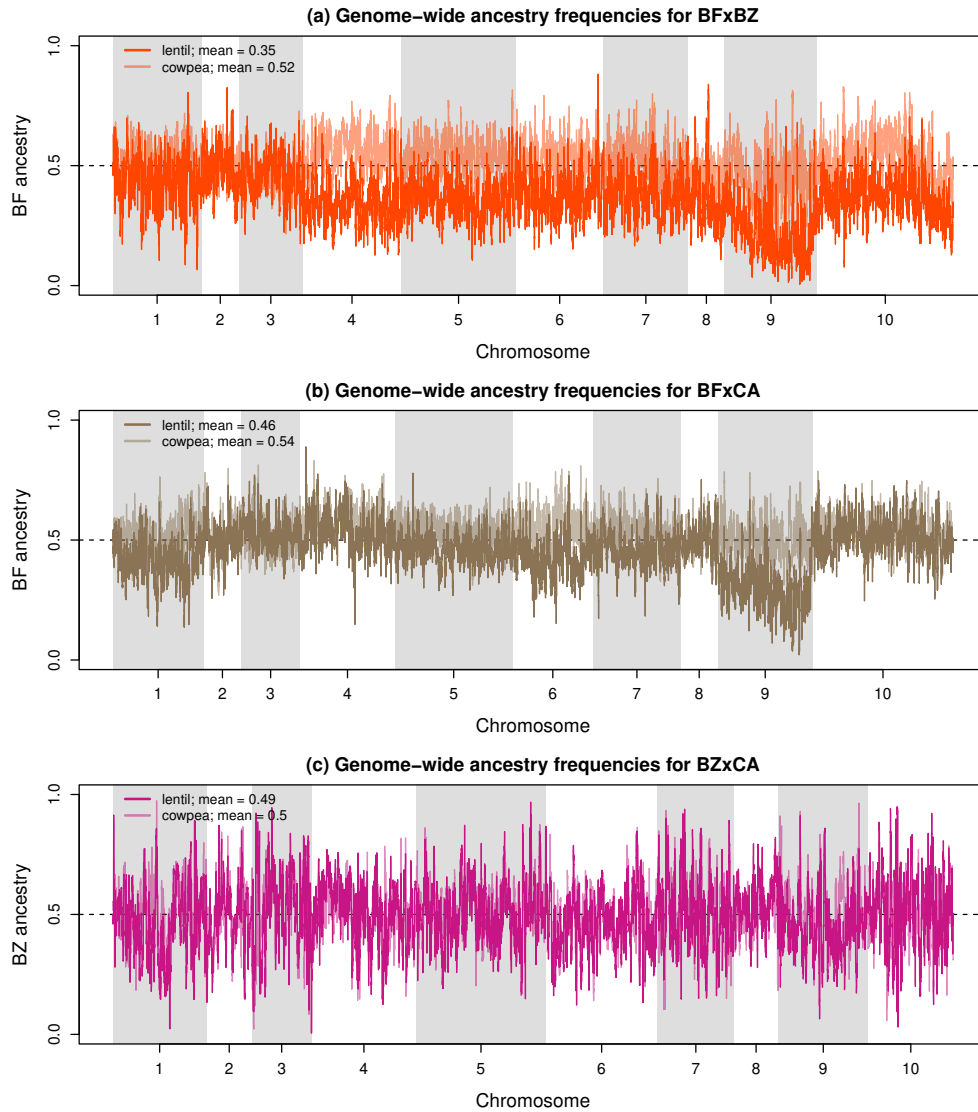


Figure 4: Genome-wide patterns of local ancestry in admixed lineages on lentil. Plots show the frequency of genetic regions inherited from one of two source populations along the genome in the admixed lines at the end of the experiment (20 generations). Lines denote averages across replicate populations with different colors (shades) for the replicate lines on lentil versus cowpea. Genome-average (mean) ancestry frequencies are also reported. Abbreviations used are: BF = Burkina Faso, BZ = Brazil and CA = California.

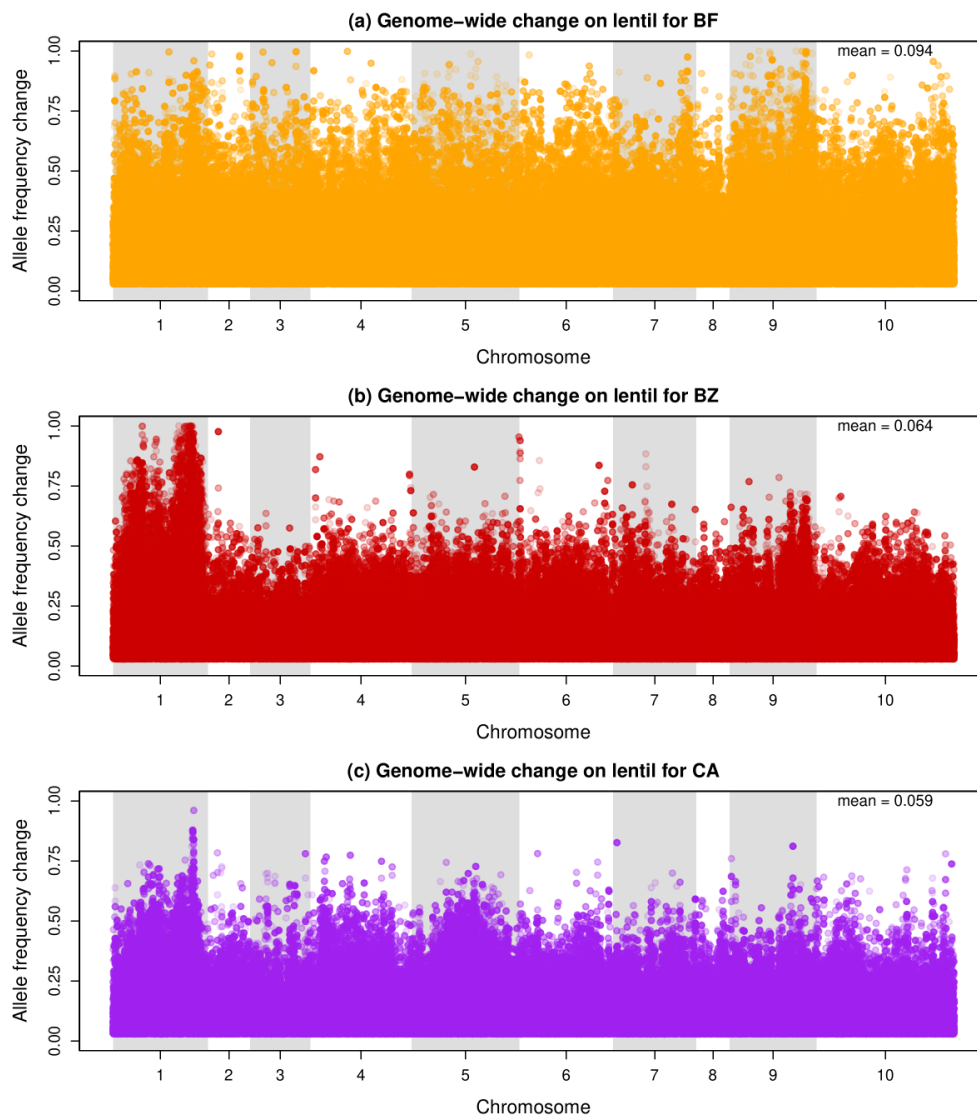


Figure 5: Manhattan plots depicting genome-wide allele frequency change for each of the non-admixed lentil-adapted lines. Results are shown for (a) Burkina Faso = BF, (b) Brazil = BZ, and (c) California = CA at the end of the experiment (after 20 generations). Points denote the unsigned (absolute) allele frequency change for each SNP, arranged in order along the 10 *C. maculatus* chromosomes. Chromosome 10 is the X chromosome. Different color shades are used for each of the five (or four for BF) replicate lines. SNPs with change  $< 0.03$  were omitted from the plot to reduce the file size. The mean change across all SNPs is reported in each panel.

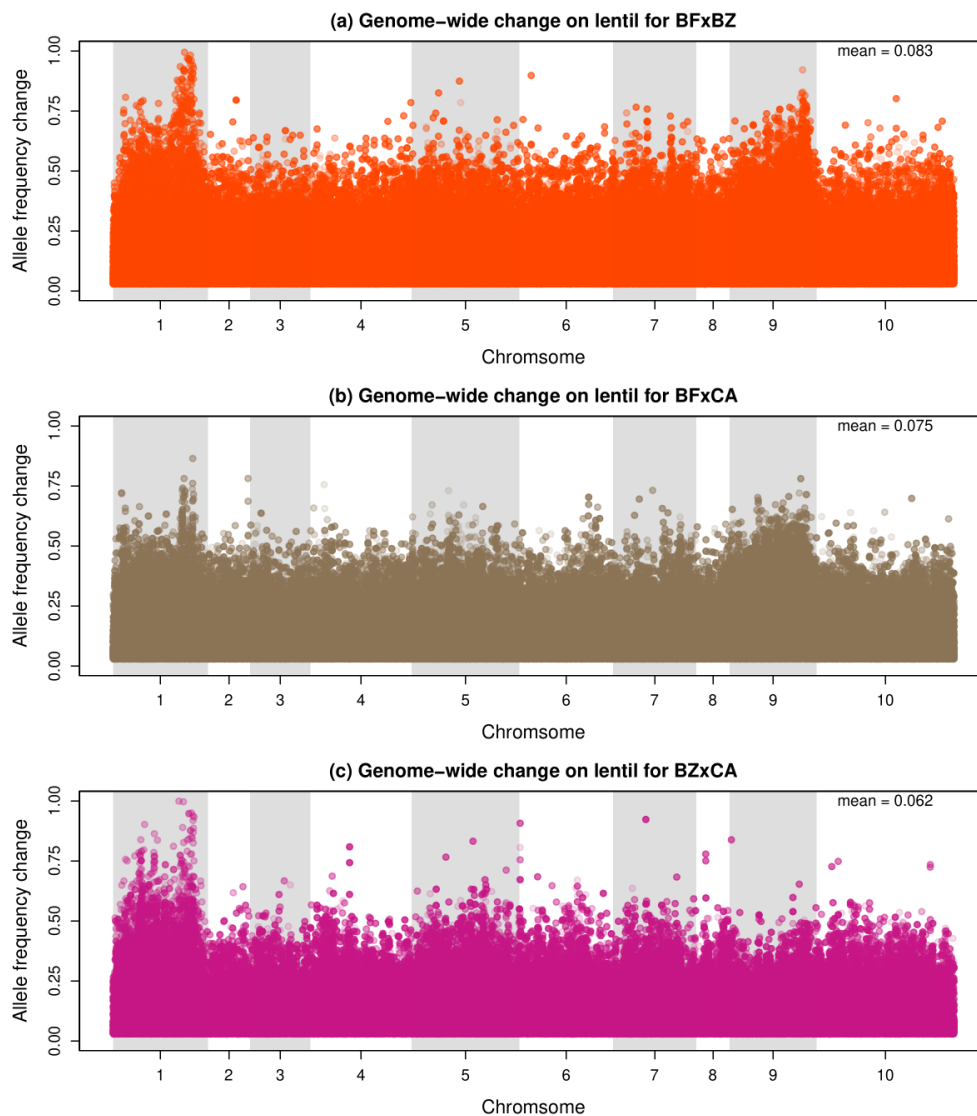


Figure 6: Manhattan plots depicting genome-wide allele frequency change for each of the admixed lentil-adapted lines. Results are shown for (a) Burkina Faso $\times$ Brazil = BF $\times$ BZ, (b) Burkina Faso $\times$ California = BF $\times$ CA, and (c) Brazil $\times$ California = BZ $\times$ CA at the end of the experiment (after 20 generations). Points denote the unsigned (absolute) allele frequency change for each SNP, arranged in order along the 10 *C. maculatus* chromosomes. Chromosome 10 is the X chromosome. Different color shades are used for each of the five (or four for BF) replicate lines. SNPs with change  $< 0.03$  were omitted from the plot to reduce the file size. The mean change across all SNPs is reported in each panel.

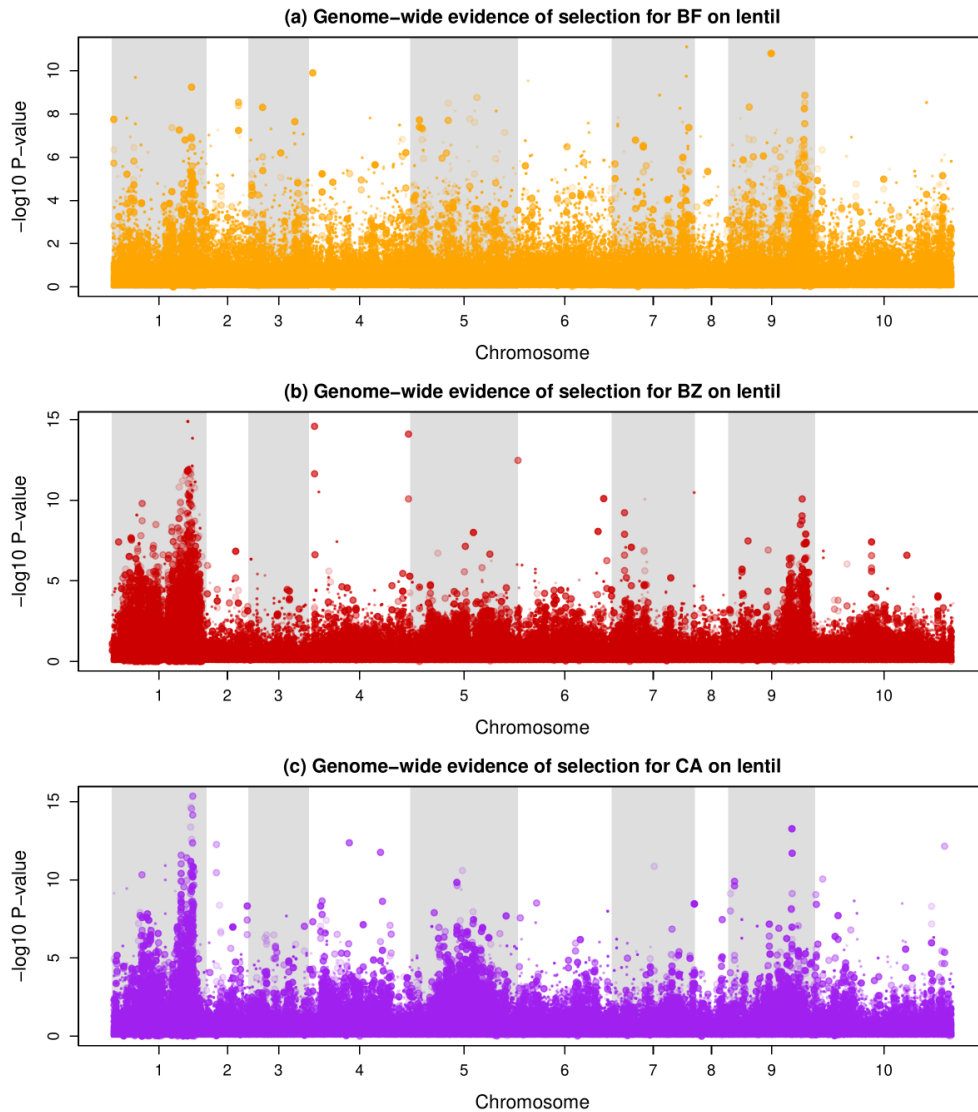


Figure 7: Manhattan plots showing evidence of allele frequency change beyond neutral expectations for each of the non-admixed lentil-adapted lines. Results are shown for (a) Burkina Faso = BF, (b) Brazil = BZ, and (c) California = CA at the end of the experiment (after 20 generations). Points denote  $-\log_{10} P$ -values from the null Wright-Fisher model for each SNP, with SNPs arranged in order along the 10 *C. maculatus* chromosomes. Chromosome 10 is the X chromosome. Different color shades are used for each of the five (or four for BF) replicate lines. Larger points are used for SNPs with significant evidence of repeated (across replicates) change beyond neutral expectations, that is,  $P$ -values from  $\text{picmin} < 0.05$  following false-discovery rate adjustment.



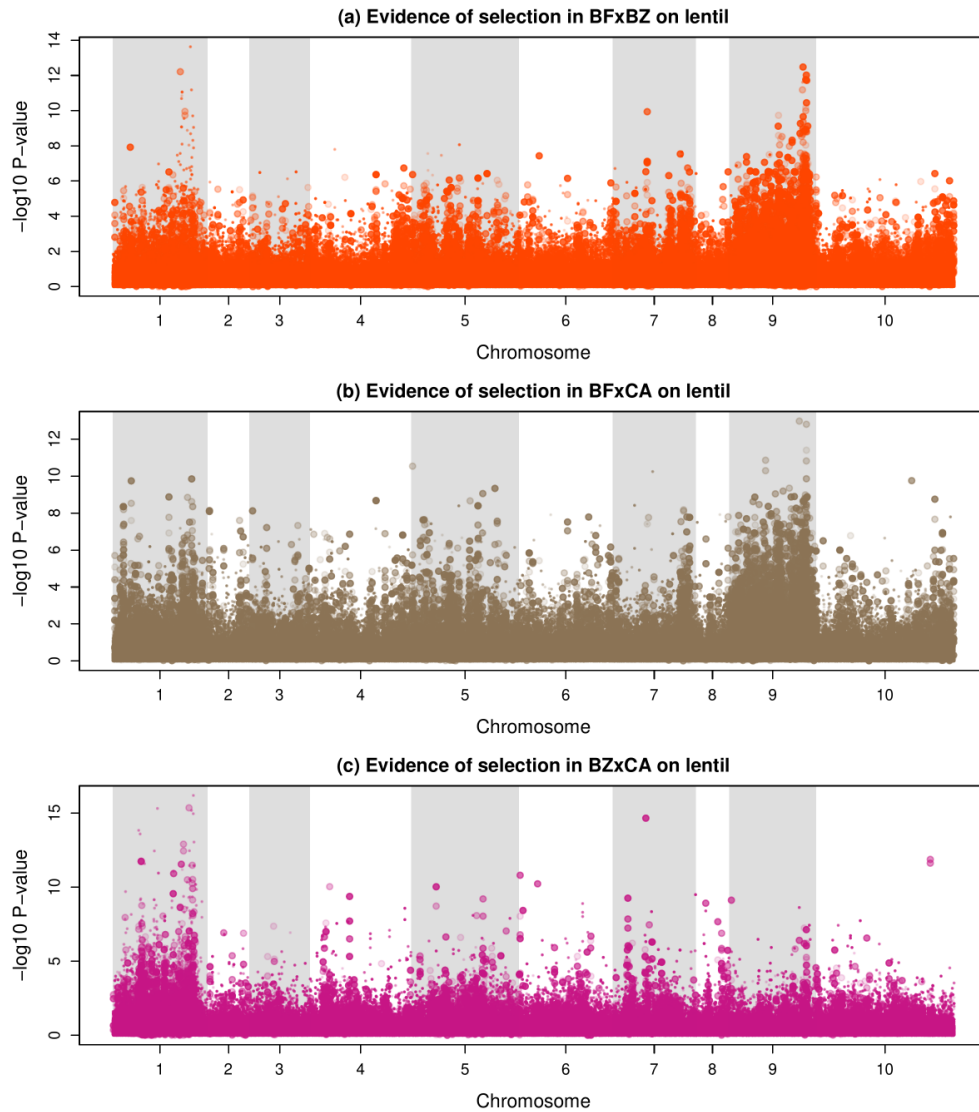


Figure 8: Manhattan plots showing evidence of allele frequency change beyond neutral expectations for each of the admixed lentil-adapted lines. Results are shown for (a) Burkina Faso $\times$ Brazil = BF $\times$ BZ, (b) Burkina Faso $\times$ California = BF $\times$ CA, and (c) Brazil $\times$ California = BZ $\times$ CA at the end of the experiment (after 20 generations). Points denote  $-\log_{10} P$ -values from the null Wright-Fisher model for each SNP, with SNPs arranged in order along the 10 *C. maculatus* chromosomes. Chromosome 10 is the X chromosome. Different color shades are used for each of the five (or four for BF) replicate lines. Larger points are used for SNPs with significant evidence of repeated (across replicates) change beyond neutral expectations, that is,  $P$ -values from `picmin`  $< 0.05$  following false-discovery rate adjustment.

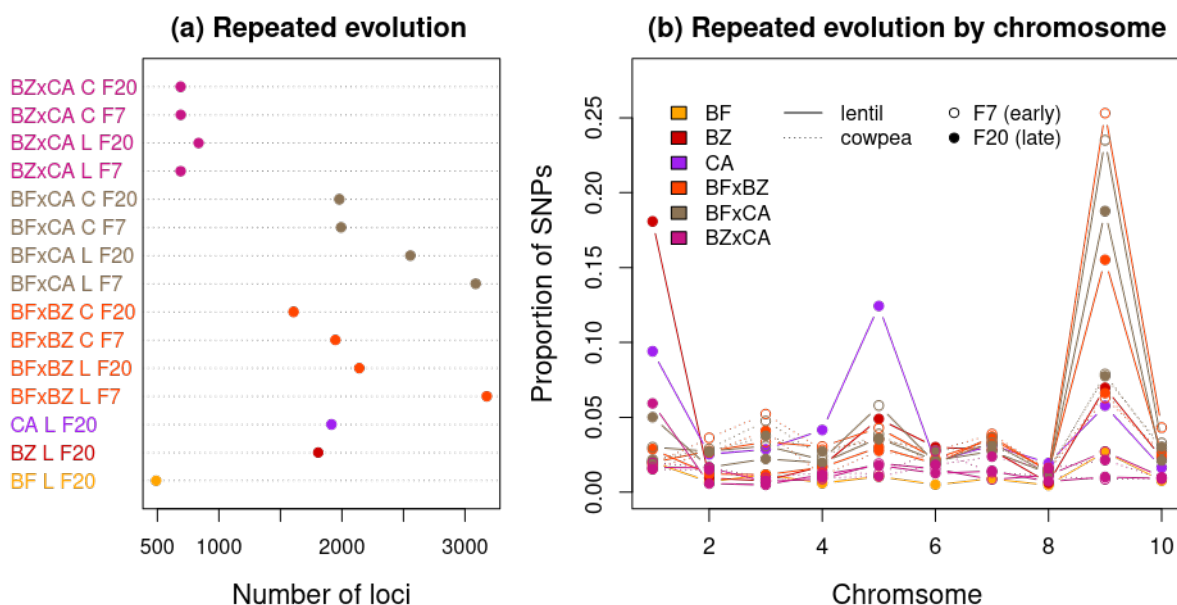


Figure 9: Graphical summary of evidence for repeated evolution among replicates from the same lineage and host treatment. Panel (a) shows the number of SNPs with significant evidence of repeated (across replicates) change beyond neutral expectations (i.e.  $P$ -values from  $\text{picmin} < 0.05$  following false-discovery rate adjustment) for each group. Abbreviations used are: BF = Burkina Faso = BF, BZ = Brazil = BZ, CA = California, C = cowpea, L = lentil, F20 = 20 generations (at the end of the experiment), and F7 = 7 generations (relatively early in the experiment). All results are based on five replicate lines except for BF L (four replicates). Panel (b) summarizes the same  $\text{picmin}$  results for each chromosome. Colored points denote the proportion of SNPs on each chromosome with significant evidence of repeated evolution; colors, point types and line types denote different experimental groups.

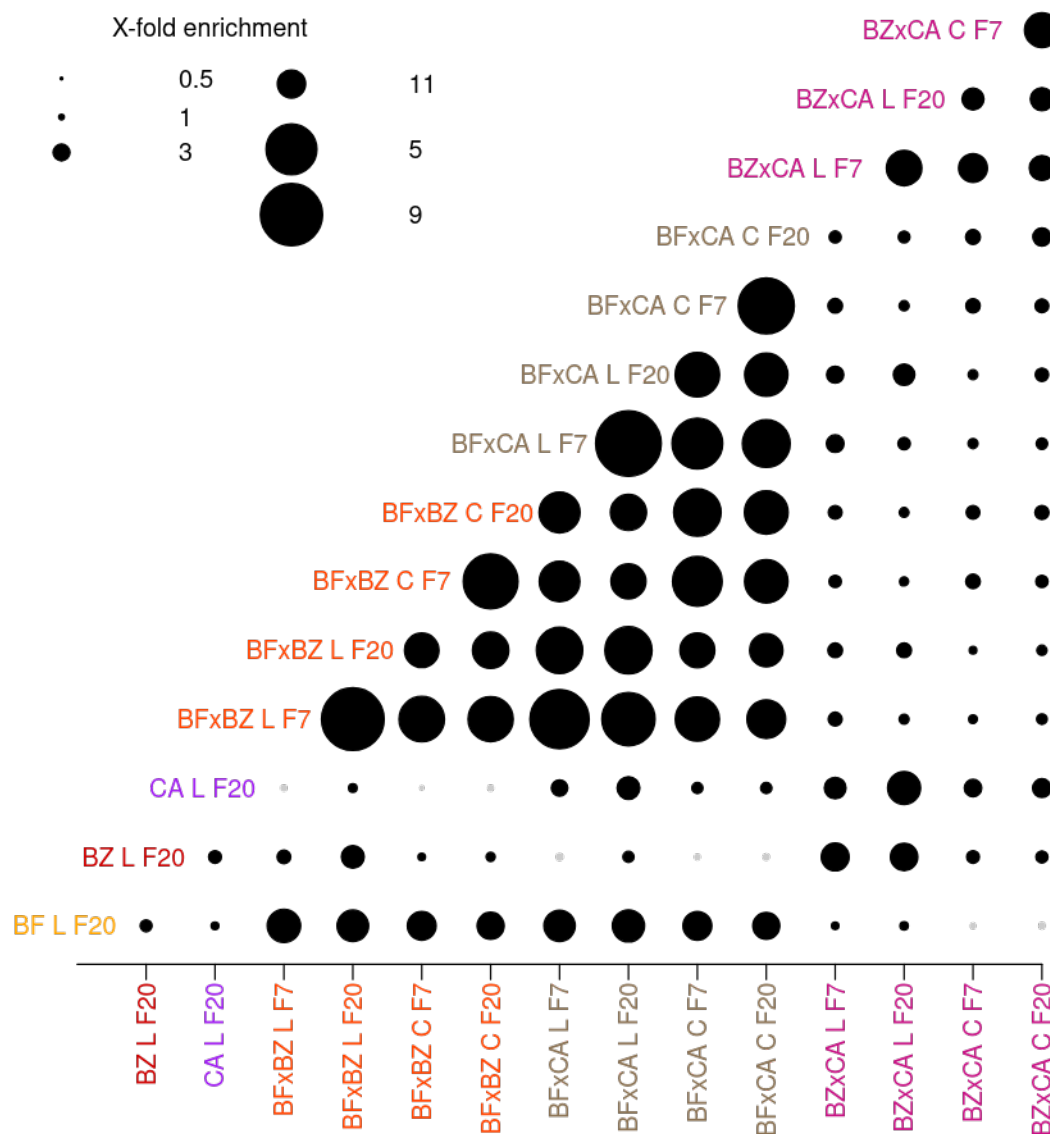


Figure 10: Graphical summary of evidence for repeated evolution between different pairs of experimental groups or time points. Each point denotes the ratio of the observed to expected number of SNPs that were among the top 5% with the strongest evidence of repeated evolution for a pair of experimental groups or time points. Larger points indicate more overlap relative to null expectations of independence between treatments (the null expectation is a 1:1 ratio). Black versus gray circles indicate pairs with ratios that are versus are not significantly greater than 1 with  $P < 0.05$  from a randomization test. Abbreviations used are: BF = Burkina Faso = BF, BZ = Brazil = BZ, CA = California, and C = cowpea, L = lentil.